

## Symmetrization and Rademacher Averages

Instructor: Sham Kakade

## 1 Rademacher Averages

Recall that we are interested in bounding the difference between empirical and true expectations uniformly over some function class  $\mathcal{G}$ . In the context of classification or regression, we are typically interested in a class  $\mathcal{G}$  that is the *loss class* associated with some function class  $\mathcal{F}$ . That is, given a *bounded* loss function  $\ell : \mathcal{D} \times \mathcal{Y} \rightarrow [0, 1]$ , we consider the class

$$\ell_{\mathcal{F}} := \{(x, y) \mapsto \ell(f(x), y) \mid f \in \mathcal{F}\} .$$

Rademacher averages give us a powerful tool to obtain uniform convergence results. We begin by examining the quantity

$$\mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \mathbb{E}[g(Z)] - \frac{1}{m} \sum_{i=1}^m g(Z_i) \right) \right] ,$$

where  $Z, \{Z_i\}_{i=1}^m$  are i.i.d. random variables taking values in some space  $\mathcal{Z}$  and  $\mathcal{G} \subseteq [a, b]^{\mathcal{Z}}$  is a set of bounded functions. We will later show that the random quantity we are interested in, namely

$$\sup_{g \in \mathcal{G}} \left( \mathbb{E}[g(Z)] - \frac{1}{m} \sum_{i=1}^m g(Z_i) \right) ,$$

will be close to the above expectation with high probability.

Let  $\epsilon_1, \dots, \epsilon_m$  be i.i.d.  $\{\pm 1\}$ -valued random variables with  $\mathbb{P}(\epsilon_i = +1) = \mathbb{P}(\epsilon_i = -1) = 1/2$ . These are also independent of the sample  $Z_1, \dots, Z_m$ . Define the *empirical Rademacher average* of  $\mathcal{G}$  as

$$\hat{\mathfrak{R}}_m(\mathcal{G}) := \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \frac{1}{m} \sum_{i=1}^m \epsilon_i g(Z_i) \mid Z_1^m \right] .$$

The *Rademacher average* of  $\mathcal{G}$  is defined as

$$\mathfrak{R}_m(\mathcal{G}) := \mathbb{E} \left[ \hat{\mathfrak{R}}_m(\mathcal{G}) \right] .$$

**Theorem 1.1.** *We have,*

$$\mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \mathbb{E}[g(Z)] - \frac{1}{m} \sum_{i=1}^m g(Z_i) \right) \right] \leq 2\mathfrak{R}_m(\mathcal{G}) .$$

*Proof.* Introduce the *ghost sample*  $Z'_1, \dots, Z'_m$ . By that we mean that  $Z'_i$ 's are independent of each other and of  $Z_i$ 's

and have the same distribution as the latter. Then we have,

$$\begin{aligned}
& \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \mathbb{E}[g(Z)] - \frac{1}{m} \sum_{i=1}^m g(Z_i) \right) \right] \\
&= \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \frac{1}{m} \sum_{i=1}^m (\mathbb{E}[g(Z)] - g(Z_i)) \right) \right] \\
&= \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \frac{1}{m} \sum_{i=1}^m \mathbb{E}[g(Z'_i) - g(Z_i) | Z_1^m] \right) \right] \\
&\leq \mathbb{E} \left[ \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \frac{1}{m} \sum_{i=1}^m (g(Z'_i) - g(Z_i)) \right) \middle| Z_1^m \right] \right] \\
&= \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \frac{1}{m} \sum_{i=1}^m (g(Z'_i) - g(Z_i)) \right) \right] \\
&= \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \frac{1}{m} \sum_{i=1}^m \epsilon_i (g(Z'_i) - g(Z_i)) \right) \right] \\
&\leq \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \frac{1}{m} \sum_{i=1}^m \epsilon_i g(Z'_i) \right] + \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \frac{1}{m} \sum_{i=1}^m \epsilon_i g(Z_i) \right] \\
&= 2\mathfrak{R}_m(\mathcal{G}) .
\end{aligned}$$

□

Since  $\mathfrak{R}_m(-\mathcal{G}) = \mathfrak{R}_m(\mathcal{G})$ , we have the following corollary.

**Corollary 1.2.** *We have,*

$$\mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \frac{1}{m} \sum_{i=1}^m g(Z_i) - \mathbb{E}[g(Z)] \right) \right] \leq 2\mathfrak{R}_m(\mathcal{G}) .$$

Since  $g(X_i) \in [a, b]$ ,

$$\sup_{g \in \mathcal{G}} \left( \mathbb{E}[g(Z)] - \frac{1}{m} \sum_{i=1}^m g(Z_i) \right)$$

does not change by more than  $(b - a)/m$  if some  $Z_i$  is changed to  $Z'_i$ . Applying the bounded differences inequality, we get the following corollary.

**Corollary 1.3.** *With probability at least  $1 - \delta$ ,*

$$\sup_{g \in \mathcal{G}} \left( \mathbb{E}[g(Z)] - \frac{1}{m} \sum_{i=1}^m g(Z_i) \right) \leq 2\mathfrak{R}_m(\mathcal{G}) + (b - a) \sqrt{\frac{\ln(1/\delta)}{2m}}$$

Recall that we denote the empirical  $\ell$ -loss minimizer by  $\hat{f}_\ell^*$ . We refer to  $L_\ell(\hat{f}_\ell^*) - \min_{f \in \mathcal{F}} L_\ell(f)$  as the estimation error. The next theorem bounds the estimation error using Rademacher averages.

## 2 Expected Regret

Now let us examine the expected regret of the empirical risk minimizer (e.g. analogous to the statistical risk). Let

$$\hat{g} = \arg \min_{g \in \mathcal{G}} \frac{1}{m} \sum_{i=1}^m g(Z_i)$$

where  $\tau$  is the training set and

$$g^* = \arg \min_{g \in \mathcal{G}} \mathbb{E} [g(Z)]$$

which is true minimizer.

**Lemma 2.1.** *The expected regret is:*

$$\begin{aligned} \mathbb{E} [\mathbb{E} [\hat{g}(Z)] - \mathbb{E} [g^*(Z)]] &\leq 2\mathfrak{R}_m(\mathcal{G}) + \mathbb{E} \left[ \frac{1}{m} \sum_{i=1}^m g^*(Z_i) - \mathbb{E} [g^*(Z)] \right] \\ &\leq 4\mathfrak{R}_m(\mathcal{G}) \end{aligned}$$

where the expectation is with respect  $\hat{g}$  (due to randomness in the training set).

*Proof.* Let

$$\hat{g}$$

The expected regret is:

$$\begin{aligned} \mathbb{E} [\mathbb{E} [\hat{g}(Z)] - \mathbb{E} [g^*(Z)]] &\leq \mathbb{E} \left[ \mathbb{E} [\hat{g}(Z)] - \frac{1}{m} \sum_{i=1}^m \hat{g}(Z_i) + \frac{1}{m} \sum_{i=1}^m \hat{g}(Z_i) - \mathbb{E} [g^*(Z)] \right] \\ &\leq \mathbb{E} \left[ \mathbb{E} [\hat{g}(Z)] - \frac{1}{m} \sum_{i=1}^m \hat{g}(Z_i) + \frac{1}{m} \sum_{i=1}^m g^*(Z_i) - \mathbb{E} [g^*(Z)] \right] \\ &\leq \mathbb{E} \left[ \sup_{g \in \mathcal{G}} \left( \mathbb{E} [\hat{g}(Z)] - \frac{1}{m} \sum_{i=1}^m \hat{g}(Z_i) \right) \right] + \mathbb{E} \left[ \frac{1}{m} \sum_{i=1}^m g^*(Z_i) - \mathbb{E} [g^*(Z)] \right] \\ &\leq 2\mathfrak{R}_m(\mathcal{G}) + \mathbb{E} \left[ \frac{1}{m} \sum_{i=1}^m g^*(Z_i) - \mathbb{E} [g^*(Z)] \right] \end{aligned}$$

The final claim is straightforward. □

## 3 Growth function

Consider the case  $\mathcal{Y} = \{\pm 1\}$  (classification). Let  $\ell$  be the 0-1 loss function and  $\mathcal{F}$  be a class of  $\pm 1$ -valued functions. We can relate the Rademacher average of  $\ell_{\mathcal{F}}$  to that of  $\mathcal{F}$  as follows.

**Lemma 3.1.** *Suppose  $\mathcal{F} \subseteq \{\pm 1\}^{\mathcal{X}}$  and let  $\ell(y', y) = \mathbf{1}[y' \neq y]$  be the 0-1 loss function. Then we have,*

$$\mathfrak{R}_m(\ell_{\mathcal{F}}) = \frac{1}{2} \mathfrak{R}_m(\mathcal{F}) .$$

*Proof.* Note that we can write  $\ell(y', y)$  as  $(1 - yy')/2$ . Then we have,

$$\begin{aligned}\mathfrak{R}_m(\ell_{\mathcal{F}}) &= \mathbb{E} \left[ \sup_{f \in \mathcal{F}} \frac{1}{m} \sum_{i=1}^m \epsilon_i \frac{1 - Y_i f(X_i)}{2} \middle| X_1^m, Y_1^m \right] \\ &= \mathbb{E} \left[ \sup_{f \in \mathcal{F}} \frac{1}{m} \sum_{i=1}^m \epsilon_i \frac{Y_i f(X_i)}{2} \middle| X_1^m, Y_1^m \right] \tag{1}\end{aligned}$$

$$\begin{aligned}&= \frac{1}{2} \mathbb{E} \left[ \sup_{f \in \mathcal{F}} \frac{1}{m} \sum_{i=1}^m (-\epsilon_i Y_i) f(X_i) \middle| X_1^m, Y_1^m \right] \\ &= \frac{1}{2} \mathbb{E} \left[ \sup_{f \in \mathcal{F}} \frac{1}{m} \sum_{i=1}^m \epsilon_i f(X_i) \middle| X_1^m, Y_1^m \right] \tag{2} \\ &= \frac{1}{2} \mathfrak{R}_m(\mathcal{F}) .\end{aligned}$$

Equation (1) follows because  $\mathbb{E}[\epsilon_i | X_1^m, Y_1^m] = 0$ . Equation (2) follows because  $-\epsilon_i Y_i$ 's jointly have the same distribution as  $\epsilon_i$ 's.  $\square$

Note that the Rademacher average of the class  $\mathcal{F}$  can also be written as

$$\mathfrak{R}_m(\mathcal{F}) = \mathbb{E} \left[ \sup_{a \in \mathcal{F}|_{X_1^m}} \frac{1}{m} \sum_{i=1}^m \epsilon_i a_i \right] ,$$

where  $\mathcal{F}|_{X_1^m}$  is the function class  $\mathcal{F}$  restricted to the set  $X_1, \dots, X_m$ . That is,

$$\mathcal{F}|_{X_1^m} := \{((f(X_1), \dots, f(X_m)) | f \in \mathcal{F})\} .$$

Note that  $\mathcal{F}|_{X_1^m}$  is finite and

$$|\mathcal{F}|_{X_1^m}| \leq \min\{|\mathcal{F}|, 2^m\} .$$

Thus we can define the *growth function* as

$$\Pi_{\mathcal{F}}(m) := \max_{x_1^m \in \mathcal{X}^m} |\mathcal{F}|_{x_1^m}| .$$

The following lemma due to Massart allows us to bound the Rademacher average in terms of the growth function.

**Lemma 3.2.** (*Finite Class Lemma*) *Let  $\mathcal{A}$  be some finite subset of  $\mathbb{R}^m$  and  $\epsilon_1, \dots, \epsilon_m$  be independent Rademacher random variables. Let  $r = \sup_{a \in \mathcal{A}} \|a\|$ . Then, we have,*

$$\mathbb{E} \left[ \sup_{a \in \mathcal{A}} \frac{1}{m} \sum_{i=1}^m \epsilon_i a_i \right] \leq \frac{r \sqrt{2 \ln |\mathcal{A}|}}{m} .$$

*Proof.* Let

$$\mu = \mathbb{E} \left[ \sup_{a \in \mathcal{A}} \sum_{i=1}^m \epsilon_i a_i \right] .$$

We have, for any  $\lambda > 0$ ,

$$\begin{aligned}
e^{\lambda\mu} &\leq \mathbb{E} \left[ \exp \left( \lambda \sup_{a \in \mathcal{A}} \sum_{i=1}^m \epsilon_i a_i \right) \right] && \text{Jensen's inequality} \\
&= \mathbb{E} \left[ \sup_{a \in \mathcal{A}} \exp \left( \lambda \sum_{i=1}^m \epsilon_i a_i \right) \right] \\
&\leq \mathbb{E} \left[ \sum_{a \in \mathcal{A}} \exp \left( \lambda \sum_{i=1}^m \epsilon_i a_i \right) \right] \\
&= \sum_{a \in \mathcal{A}} \mathbb{E} \left[ \exp \left( \lambda \sum_{i=1}^m \epsilon_i a_i \right) \right] \\
&= \sum_{a \in \mathcal{A}} \prod_{i=1}^m \mathbb{E} [\exp (\lambda \epsilon_i a_i)] \\
&\leq \sum_{a \in \mathcal{A}} \prod_{i=1}^m e^{\lambda^2 a_i^2 / 2} && \because \text{Hoeffding's lemma} \\
&= \sum_{a \in \mathcal{A}} e^{\lambda^2 \|a\|^2 / 2} \\
&\leq |\mathcal{A}| e^{\lambda^2 r^2 / 2}
\end{aligned}$$

Taking logs and dividing by  $\lambda$ , we get that, for any  $\lambda > 0$ ,

$$\mu \leq \frac{\ln |\mathcal{A}|}{\lambda} + \frac{\lambda r^2}{2}.$$

Setting  $\lambda = \sqrt{2 \ln |\mathcal{A}| / r^2}$  gives,

$$\mu \leq r \sqrt{2 \ln |\mathcal{A}|},$$

which proves the lemma. □