

CSE599s Counting and Sampling: Homework 1

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Problem 1 Solution

Let $D_{xy}(t) = \|K^t(x, \cdot) - K^t(y, \cdot)\|_{TV}$ and $D(t) = \max_{x, y \in \Omega} D_{xy}(t)$. For every t , since $\pi = \sum_{x \in \Omega} \pi(x) K^t(x, \cdot)$ is a convex combination of $K^t(x, \cdot)$ for $x \in \Omega$. We know that,

$$\Delta(t) = \max_{x \in \Omega} \|K^t(x, \cdot) - \pi\|_{TV} \leq \max_{x, y \in \Omega} \|K^t(x, \cdot) - K^t(y, \cdot)\|_{TV} = D(t).$$

On the other hand, we know that

$$D(t) = \max_{x, y \in \Omega} \|K^t(x, \cdot) - K^t(y, \cdot)\|_{TV} \leq 2 \max_{x \in \Omega} \|K^t(x, \cdot) - \pi\|_{TV} = 2\Delta(t).$$

Let $X_0 = x$ and $Y_0 = y$. By the coupling lemma, there exists a coupling of $K^t(x, \cdot)$ and $K^t(y, \cdot)$ so that

$$\Pr[X_t \neq Y_t] = \|K^t(x, \cdot) - K^t(y, \cdot)\|_{TV} = D_{xy}(t).$$

Then for $s \geq 1$, we construct a coupling of X_{t+s} and Y_{t+s} as follows:

- If $X_t = Y_t$, then $X_{t+i} = Y_{t+i}$ for $i = 1, \dots, s$.
- If $X_t = x'$ and $Y_t = y'$ where $x' \neq y'$, then by the coupling lemma we are able to construct X_{t+s} and Y_{t+s} so that

$$\Pr[X_{t+s} \neq Y_{t+s} | X_t = x', Y_t = y'] = \|K^s(x', \cdot) - K^s(y', \cdot)\|_{TV} = D_{x'y'}(s) \leq D(s).$$

By now we have that for every x, y

$$D_{xy}(t+s) \leq \Pr[X_{t+s} \neq Y_{t+s}] \leq D(s)D_{xy}(t) \leq D(s)D(t),$$

where the second inequality hold because $X_{t+s} \neq Y_{t+s}$ only when $X_t \neq Y_t$. Since the above inequality holds for every x, y , we have

$$D(t+s) = \max_{x, y \in \Omega} D_{x,y}(t+s) \leq D(s)D(t).$$

And thus $D(kt) \leq D(t)^k$ for any k .

Finally, by the relationship between $\Delta(t)$ and $D(t)$, we have

$$\Delta(\tau_{mix} \log(1/\epsilon)) \leq D(\tau_{mix} \log(1/\epsilon)) \leq D(\tau_{mix})^{\log(1/\epsilon)} \leq (2\Delta(\tau_{mix}))^{\log(1/\epsilon)} \leq \left(\frac{1}{e}\right)^{\log(1/\epsilon)} = \epsilon.$$

Therefore, $\tau(\epsilon) \leq O(\tau_{mix} \log(1/\epsilon))$.

Reference

<https://people.eecs.berkeley.edu/~sinclair/cs294/n3.pdf>

Problem 2 Solution

We define a new sequence of random variables (Y_i) by $Y_i = X_i^2 - (1 - \theta)i$. Then,

$$\begin{aligned} & \mathbb{E}[Y_{t+1} | X_0, \dots, X_t] \\ &= \mathbb{E}[X_{t+1}^2 - (1 - \theta)(t + 1) | X_0, \dots, X_t] \\ &= \frac{1 - \theta}{2}(X_t + 1)^2 + \frac{1 - \theta}{2}(X_t - 1)^2 + \theta X_t^2 - (1 - \theta)(t + 1) \\ &= X_t^2 - (1 - \theta)t \\ &= Y_t. \end{aligned}$$

Therefore, (Y_i) is a martingale with respect to (X_i) .

For $i = -b, \dots, b$, let T_i to be the expected number of steps to first reach $+b$ or $-b$ starting from i . Then $T_b = T_{-b} = 0$ and $T_i = (\frac{1-\theta}{2}T_{i-1} + \frac{1-\theta}{2}T_{i+1} + \theta T_i) + 1$ for $i = -b+1, \dots, b-1$. It is clear that this set of equations have a finite solution. Therefore $\mathbb{E}[T] = T_0$ is bounded. On the other hand, for all t before reaching $+b$ or $-b$ we have $\mathbb{E}[|Y_{t+1} - Y_t| | X_0, \dots, X_t] \leq 2b+1 + (1-\theta)$ which is bounded. Therefore, by the optional stopping theorem, we have

$$\mathbb{E}[X_T^2 - (1 - \theta)T] = \mathbb{E}[Y_T] = \mathbb{E}[Y_0] = 0.$$

Thus we have

$$\mathbb{E}[T] = \frac{\mathbb{E}[X_T^2]}{1 - \theta} = \frac{b^2}{1 - \theta}.$$

Problem 3(a) Solution

Consider the following coupling between X_t and Y_t . Suppose that X_t and Y_t differ in d bits i_1, \dots, i_d , then with $\frac{1}{n+1}$ probability X_t remains unchanged, otherwise it uniformly flips a random bit. And,

- In all cases when X_t flips a bit on which X_t and Y_t agree, Y_t do the same.
- If d is an even number, then when X_t flips the bit i_k (on which they differ), Y_t flips the bit i_{d-k+1} . And when X_t remains unchanged, Y_t also remains unchanged.
- If d is an odd number, then for $k = 1, \dots, d - 1$ when X_t flips the bit i_k , Y_t flips the bit i_{d-k} . When X_t flips the last bit they differ, Y_t remains unchanged. And when X_t remains unchanged, Y_t flips the last bit they differ.

In the above coupling, the distance between X_t and Y_t decreases by 2 with probability $\frac{d}{n+1}$ when d is even and with probability $\frac{d+1}{n+1}$ when d is odd. Let T be the stopping time when X_T and Y_T agree. Then the expectation of T is bounded by

$$\max\left\{ \sum_{k \geq 1, 2k-1 \leq n} \frac{n+1}{2k-1}, \sum_{k \geq 1, 2k \leq n} \frac{n+1}{2k} \right\} \leq \frac{1}{2}n(\ln n + 5).$$

Since the above time is highly concentrated around its expectation (for the same reason of the coupon collector problem) in the sense that

$$\Pr[T > \frac{1}{2}n(\ln n + 5) + cn] \leq e^{-c},$$

by Lemma 4.9 in Lecture Note 4, the mixing time τ_{mix} is upper bounded by $\frac{1}{2}n \ln n + O(n)$.

Reference

<https://people.eecs.berkeley.edu/~sinclair/cs294/n5.pdf>

Problem 4 Solution

Here we define the distance between two configurations σ^1 and σ^2 as $d(\sigma^1, \sigma^2) = \sum_i \mathbf{1}_{\sigma_i^1 \neq \sigma_i^2}$. Let X_t and Y_t be two configurations that differ only on site i on the torus (i.e. $d(X_t, Y_t) = 1$). And suppose that $X_t(i) = +$ and $Y_t(i) = -$. We consider the following coupling: X^t and Y^t pick the same uniformly random site and maximize the probability of agreement. Let j be the site X_t and Y_t pick. In this coupling, we have

- If $j = i$, then $d(X_{t+1}, Y_{t+1}) = 0$ with probability 1.
- If j is a neighbor of i . Let α be the number of +'s minus the number of -'s in the neighborhood of j in X_t (or Y_t) that excludes site i . Then we have

$$\Pr[X_{t+1}(j) = +] = e^{2\beta(\alpha+1)} / (e^{2\beta(\alpha+1)} + e^{2\beta(-\alpha-1)}),$$

and

$$\Pr[Y_{t+1}(j) = +] = e^{2\beta(\alpha-1)} / (e^{2\beta(\alpha-1)} + e^{2\beta(-\alpha+1)}).$$

Therefore

$$\begin{aligned} & \Pr[d(X_{t+1}, Y_{t+1}) = 2] \\ &= \Pr[X_{t+1}(j) \neq Y_{t+1}(j)] \\ &= \Pr[X_{t+1}(j) = +] - \Pr[Y_{t+1}(j) = +] \\ &= \frac{e^{2\beta} - e^{-2\beta}}{e^{2\alpha\beta} + e^{-2\alpha\beta} + e^{2\beta} + e^{-2\beta}} \\ &\leq \frac{e^{2\beta} - e^{-2\beta}}{e^{2\beta} + e^{-2\beta} + 2}, \end{aligned}$$

where the last inequality is tight when $\alpha = 0$.

Note that in this case when $d(X_{t+1}, Y_{t+1}) \neq 2$, we always have $d(X_{t+1}, Y_{t+1}) = 1$.

- If $j \neq i$ and j is not a neighbor of i then $d(X_{t+1}, Y_{t+1}) = 1$ with probability 1.

According to the above, the expected distance between X_{t+1} and Y_{t+1} is bounded by:

$$\mathbb{E}[d(X_{t+1}, Y_{t+1})] \leq 1 + \frac{4}{n} \cdot \frac{e^{2\beta} - e^{-2\beta}}{e^{2\beta} + e^{-2\beta} + 2} - \frac{1}{n}.$$

When $\beta < 0.01$, $\frac{e^{2\beta} - e^{-2\beta}}{e^{2\beta} + e^{-2\beta} + 2} < \frac{1}{8}$, and therefore $\mathbb{E}[d(X_{t+1}, Y_{t+1})] < 1 - \frac{1}{2n}$. Since the diameter of this metric is $O(n)$, by the path coupling lemma, the mixing time is bounded by $O(n \log n)$.

Reference

<https://people.eecs.berkeley.edu/~sinclair/cs294/n20.pdf>