

Disease Association Studies

Lectures 7 – Oct 19, 2011

CSE 527 Computational Biology, Fall 2011

Instructor: Su-In Lee TA: Christopher Miles

Monday & Wednesday 12:00-1:20

Johnson Hall (JHN) 022

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Last Class ... Haplotype reconstruction genetic markers ...ACTCGG T(AT CGGACTC TT G TCGGCCCG rcggc(cc cg CCGGACC CGGGACT CGGTAGGCCT T TCGGCCGC ...ACCCGGTTGGCCT ...ACCCGGTT(CGGGACC CC G TT C CGGCCGG CCGGANCO CC GCCTATATTCGGCCCC CC ...ACCCGGTA ...ACTCGG A C<mark>G</mark>GGACCCGGTTGGCCTT TATTCGGCCCG GCCTATATTCGGCCGGC ...ACTCGG^TA(C**G**GGA<mark>CT</mark> ...ACTCGG TO ...ACCCGG TAC ΑT Single nucleotide polymorphism (SNP) [snip] = a variation at a single site in DNA

Outline

- Application to disease association analysis
 - Single marker based association tests
 - Haplotype-based approach
 - Indirect association predicting unobserved SNPs



- Selection of tag SNPs
- Genetic linkage analysis
 - Pedigree-based gene mapping
 - Elston-Stewart algorithm
 - Association vs linkage

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A single marker association test

- Data
 - Genotype data from case/control individuals
 - e.g. case: patients, control: healthy individuals
- Goals
 - Compare frequencies of particular alleles, or genotypes, in set of cases and controls
 - Typically, relies on standard contingency table tests
 - Chi-square goodness-of-fit test
 - Likelihood ratio test
 - Fisher's exact test

Construct contingency table

- Organize genotype counts in a simple table
 - Rows: one row for cases, another for controls
 - Columns: one of each genotype (or allele)
 - Individual cells: count of observations

i: case, control j: 0/0, 0/1, 1/1		j=1	j=2	j=3	
		0/0	0/1	1/1	
i=1	Case (affected)	O _{1,1}	O _{1,2}	O _{1,3}	$\mathbf{O}_{1, \cdot} = \mathbf{O}_{1, 1} + \mathbf{O}_{1, 2} + \mathbf{O}_{1, 3}$
i=2	Control (unaffected)	O _{2,1}	O _{2,2}	O _{2,3}	$O_{2,} = O_{2,1} + O_{2,2} + O_{2,3}$
		$0_{1} = 0_{11} + 0_{21}$	0. 2=012+022	0, 2=0, 2+0, 2	

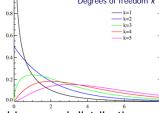
- Notation
 - Let O_{ii} denote the observed counts in each cell
 - Let E_{ij} denote the expected counts in each cell
 E_{ij} = O_{i,.} O_{.,j} / O_{.,j}.

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Goodness of fit tests (1/2)

- Null hypothesis
 - There is no statistical dependency between the genotypes and the phenotype (case/control)
- P-value
 - Probability of obtaining a test statistic at least as extreme as the one that was actually observed
- Chi-square test

$$\chi^{2} = \sum_{i,j} \frac{(O_{i,j} - E_{i,j})^{2}}{E_{i,j}}$$



- If counts are large, compare statistic to chi-squared distribution
 - p = 0.05 threshold is 5.99 for 2 df (degrees of freedom, e.g. genotype test)
 - p = 0.05 threshold is 3.84 for 1 df (e.g. allele test)
- If counts are small, exact or permutation tests are better

Goodness of fit tests (2/2)

- Likelihood ratio test
 - The test statistics (usually denoted D) is twice the difference in the log-likelihoods:

$$D = -2 \ln \left(\frac{\text{likelihood for null model}}{\text{likelihood for alternative model}} \right)$$

$$= -2 \ln \frac{\prod_{i,j} (E_{i,j}/O)^{O_{i,j}}}{\prod_{i,j} (O_{i,j}/O)^{O_{i,j}}} = 2 \sum_{i,j} O_{i,j} \ln \frac{O_{i,j}}{E_{i,j}}$$

- How about we do this for haplotypes?
 - When does it out-perform the single marker association test?

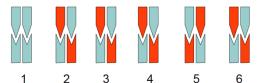
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Haplotype association tests

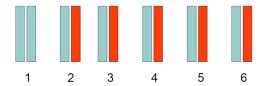
- Calculate haplotype frequencies in each group
- Find most likely haplotype for each group
- Fill in contingency table to compare haplotypes in the two groups (case, control)
- Not recommended!

Case genotypes & haplotypes

Observed case genotypes



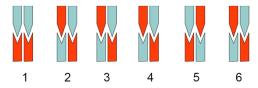
- The phase reconstruction in the five ambiguous individuals will be driven by the haplotypes observed in individual 1 ...
- Inferred case haplotypes



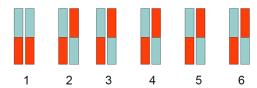
This kind of phenomenon will occur with nearly all population based haplotyping methods!

Control genotypes & haplotypes

Observed control genotypes



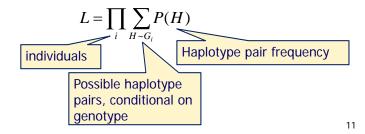
- Note these are identical, except for the single homozygous individual ...
- Inferred case haplotypes



 Oops... The difference in a single genotype in the original data has been greatly amplified by estimating haplotypes...

Haplotype association tests

- Never impute haplotypes in two groups separately
- Alternatively,
 - Consider both samples jointly
 - Schaid et al (2002) Am J Hum Genet 70:425-34
 - Zaytkin et al (2002) Hum Hered. 53:79-91
 - Use maximum likelihood



Likelihood-based test

- Calculate 3 likelihoods
 - Maximum likelihood for combined samples, L_A
 - Maximum likelihood for control sample, L_R
 - Maximum likelihood for case sample, L_C

$$D = 2\ln\left(\frac{L_B L_C}{L_A}\right) \sim \chi_{df}^2$$

 df (degrees of freedom) corresponds to number of non-zero haplotype frequencies in large samples

Significance in small samples

- In reality sample sizes, it is hard to estimate the number of df accurately
- Instead, use a permutation approach to calculate empirical significance levels
- How?

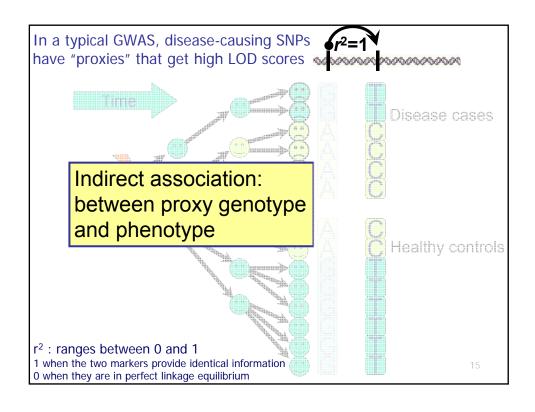
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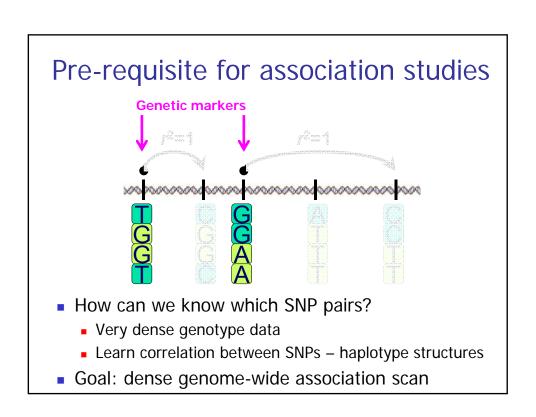
Outline

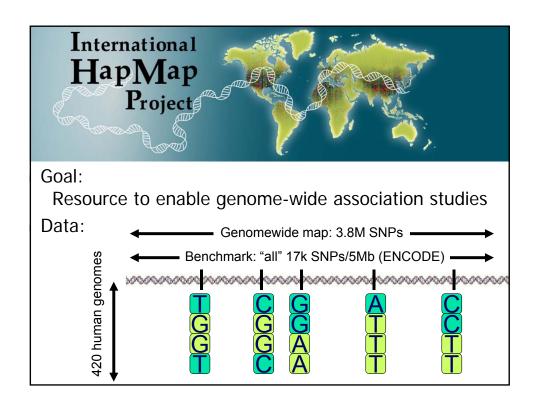
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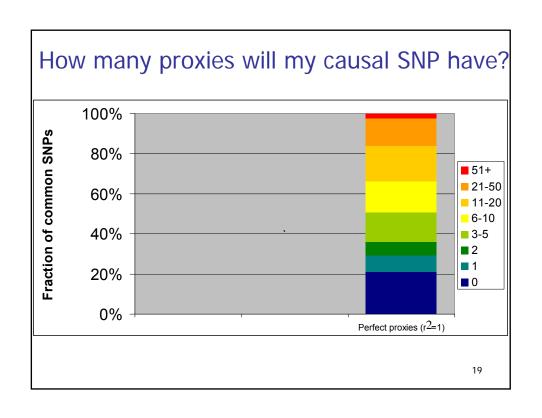
Main question for HapMap:

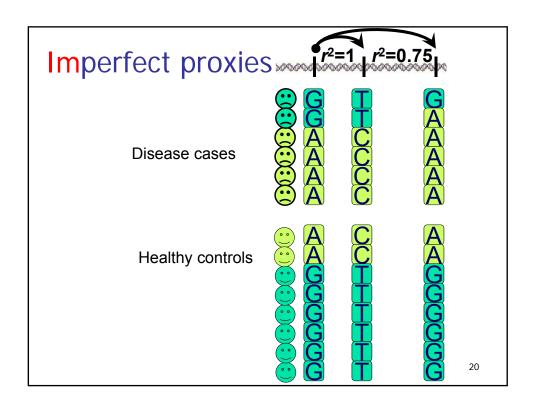


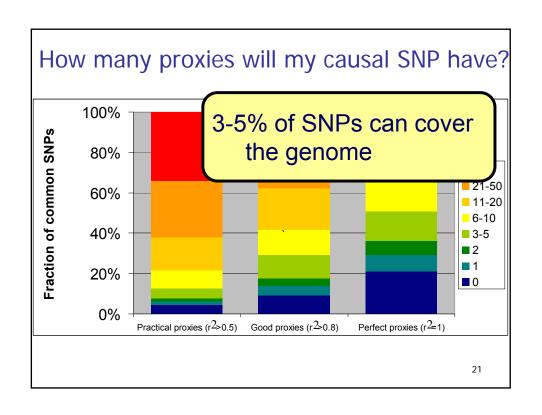
Are genomewide association studies doable?

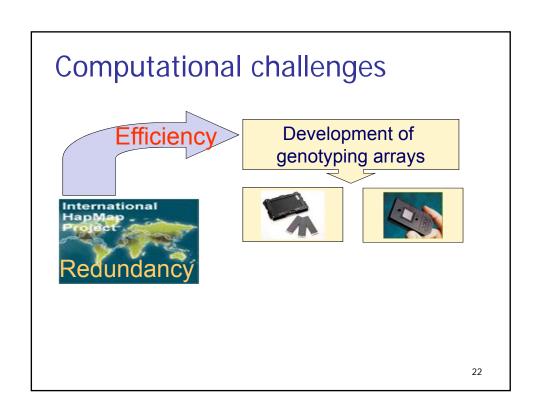
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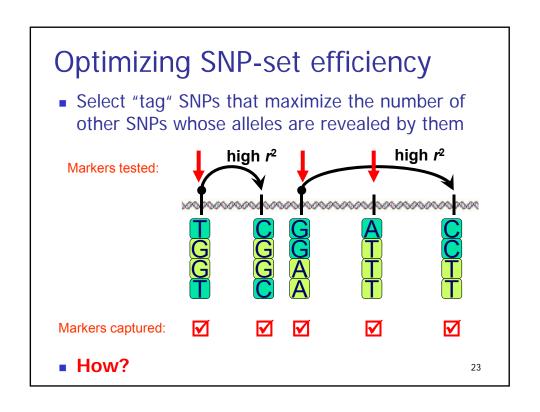
Do SNPs have enough proxies?

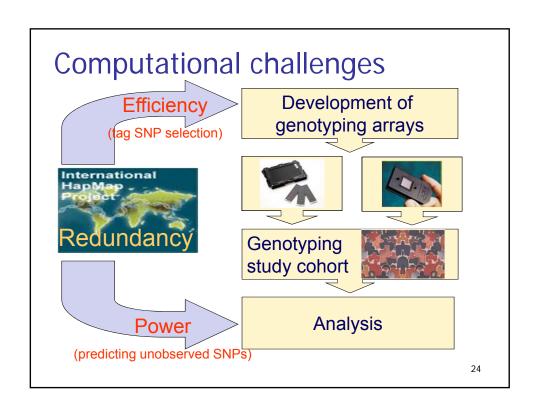








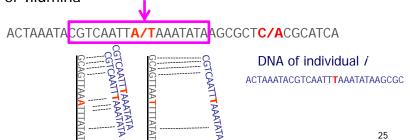


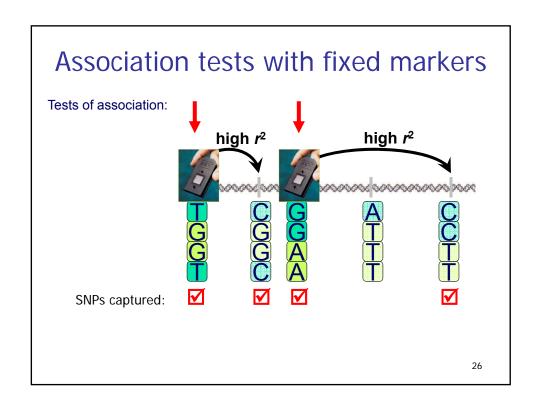


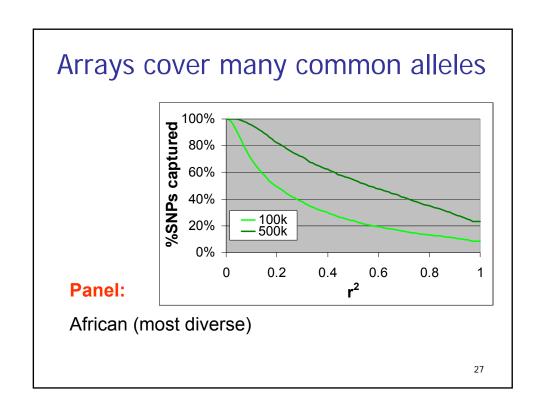
Analysis questions

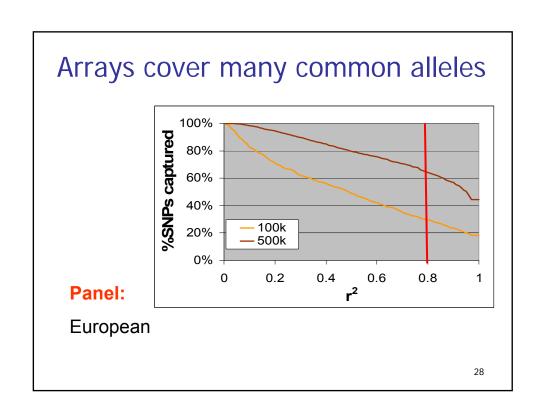


- Can we quantify the coverage of common sequence variations measured by genome-wide SNP genotyping arrays?
- SNP genotyping arrays
 - Arrays covering 100K/500K/1M SNPs from Affymetrix or Illumina





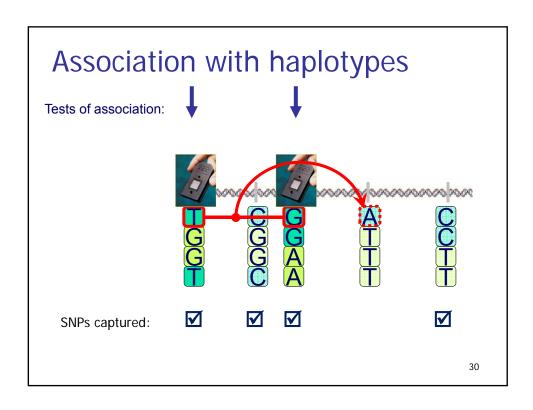


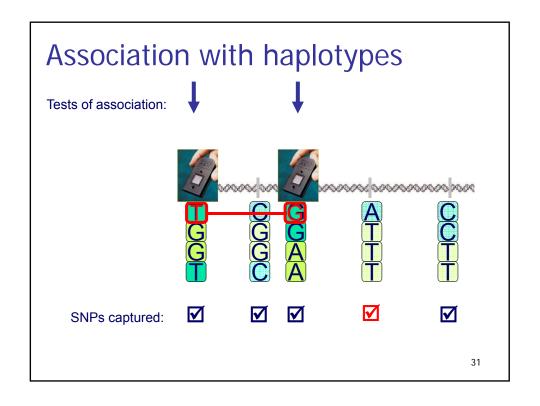


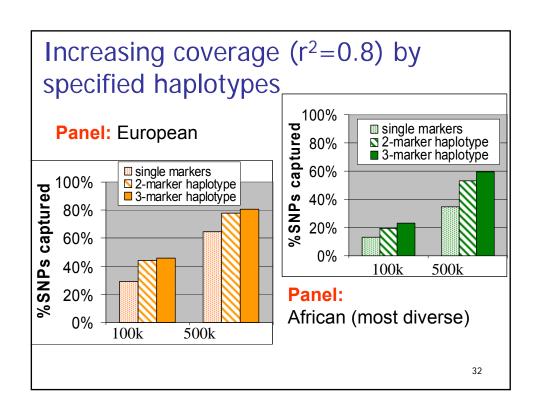
Analysis questions

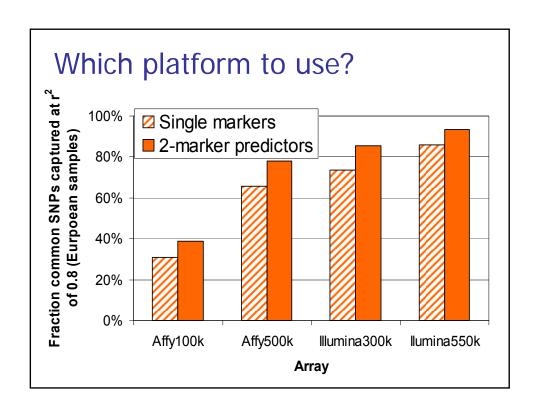


- Can we quantify the coverage of common sequence variations measured by genome-wide SNP genotyping arrays?
- Can we do better?









Summary

- Association analysis is a powerful strategy for common disease research
- HapMap and genomewide technologies enable whole-genome association scans